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Emerging bias mechanisms in sibling comparison designs

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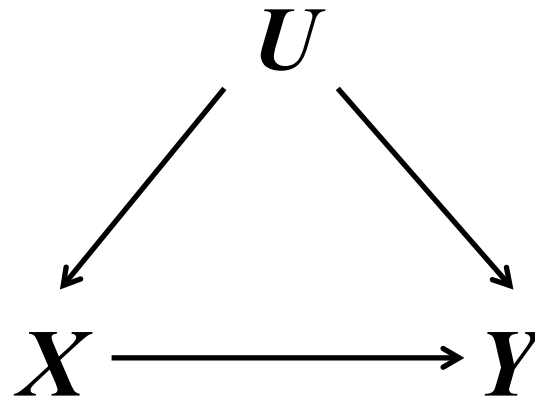
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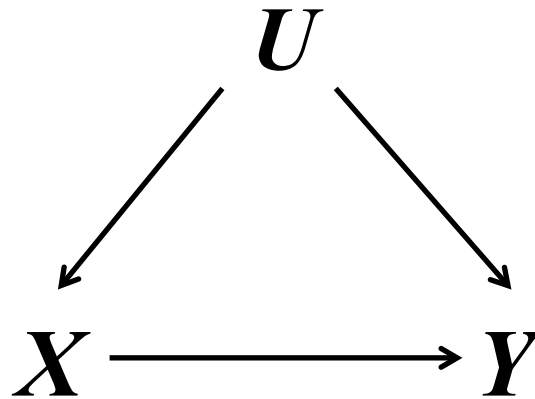
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- Frisell Thomas, Öberg Sara, Kuja-Halkola Ralf, Sjölander Arvid. **Sibling Comparison Designs: Bias From Non-Shared Confounders and Measurement Error.** *Epub ahead of print Epidemiology 2012*

Confounding DAG

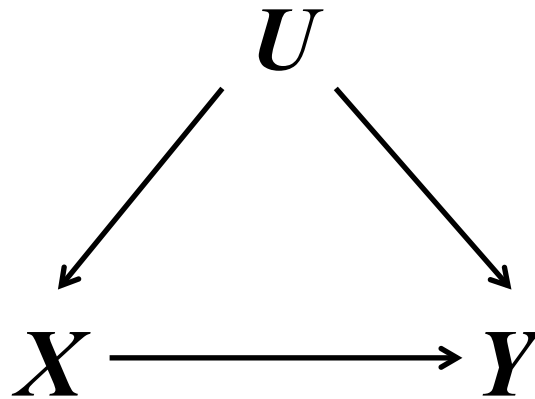


Confounding DAG



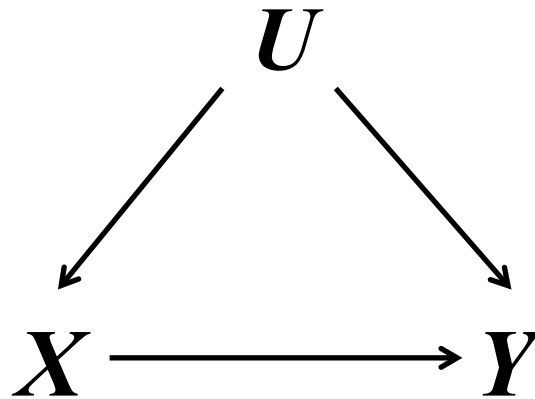
1. Randomize people to different levels of X
2. Measure U and adjust using e.g. regression
3. If U unmeasured, or even unknown, comparing siblings might be a good idea

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Sibling comparisons

- Analyses:
 - Generalized Linear Mixed Models
 - Generalized Estimation Equations
 - Conditional regressions
 - Fixed effects (in the econometric/social sciences sense)
 - **Between-Within (BW)**

Within sibling estimate

- Standard model vs. BW model

- “Cohort” estimate:

$$g(E(Y_{ij} | X_{ij})) = \mu + \beta_C X_{ij} \quad (1)$$

- Within estimate:

$$g(E(Y_{ij} | X_{ij}, X_{i\cdot})) = \mu + \beta_B X_{i\cdot} + \beta_W X_{ij} \quad (2)$$

Within sibling estimate

- Standard model vs. BW model

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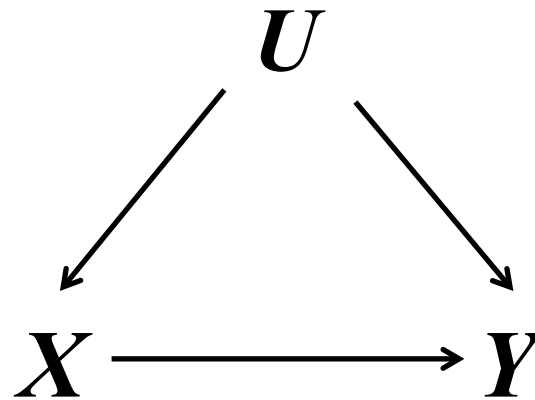
$$g(E(Y_{ij} | X_{ij}, X_{i\cdot})) = \mu + \beta_B X_{i\cdot} + \beta_W (X_{ij} - X_{i\cdot})$$

What is the problem then?

- We want to control for unmeasured confounders!
- Inference comes from exposure discordant pairs
- Isn't the exposure discordant pairs probably less similar in unmeasured confounders than a randomly picked pair of individuals with the same exposure levels?
- Correct inference?!

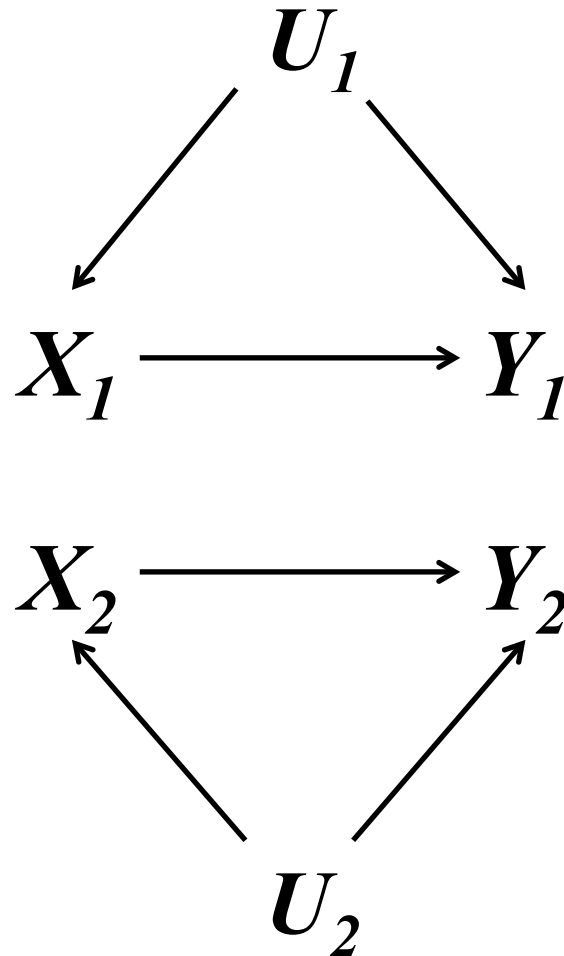
Causality

DAG

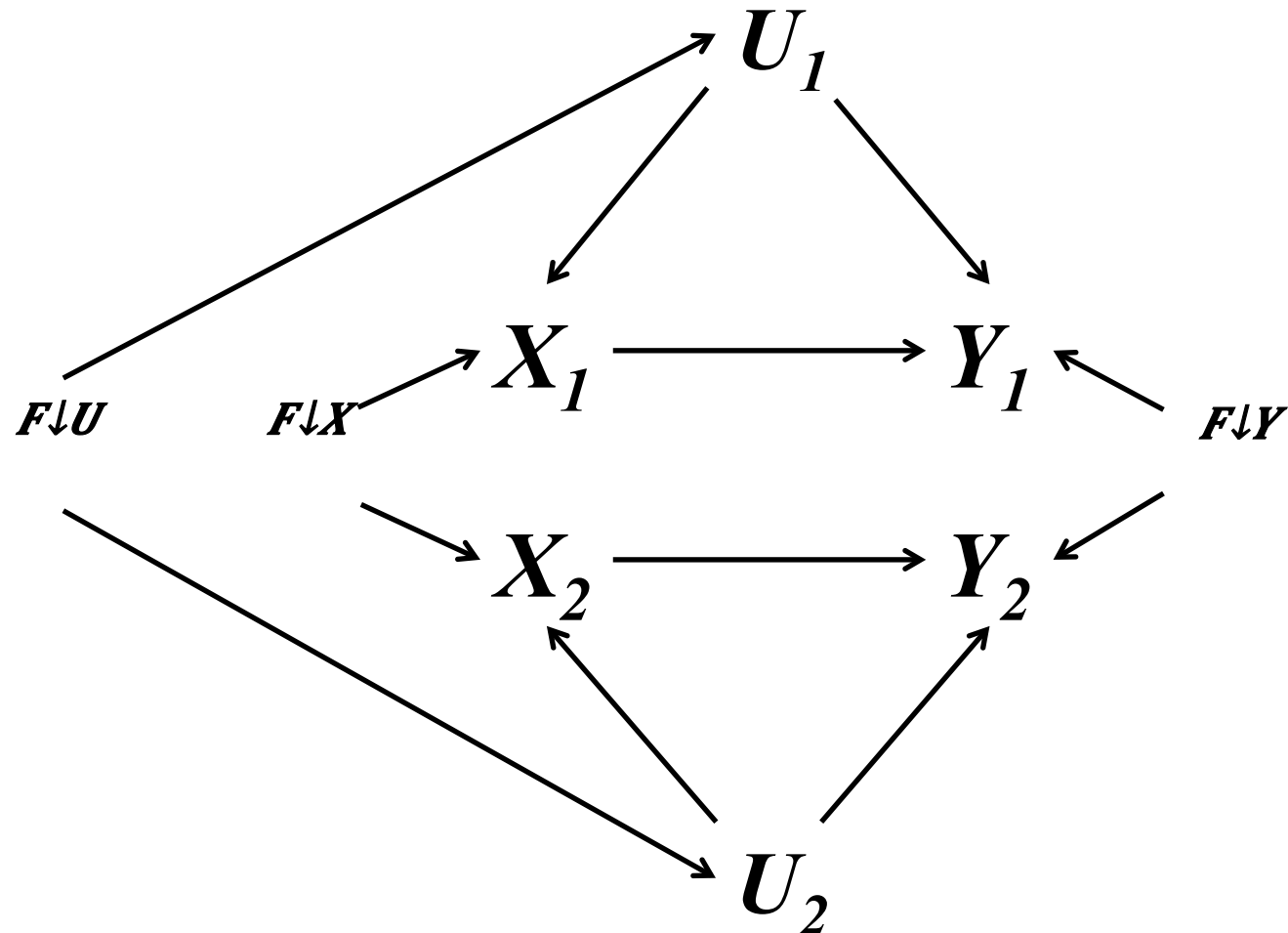


Causality

DAG

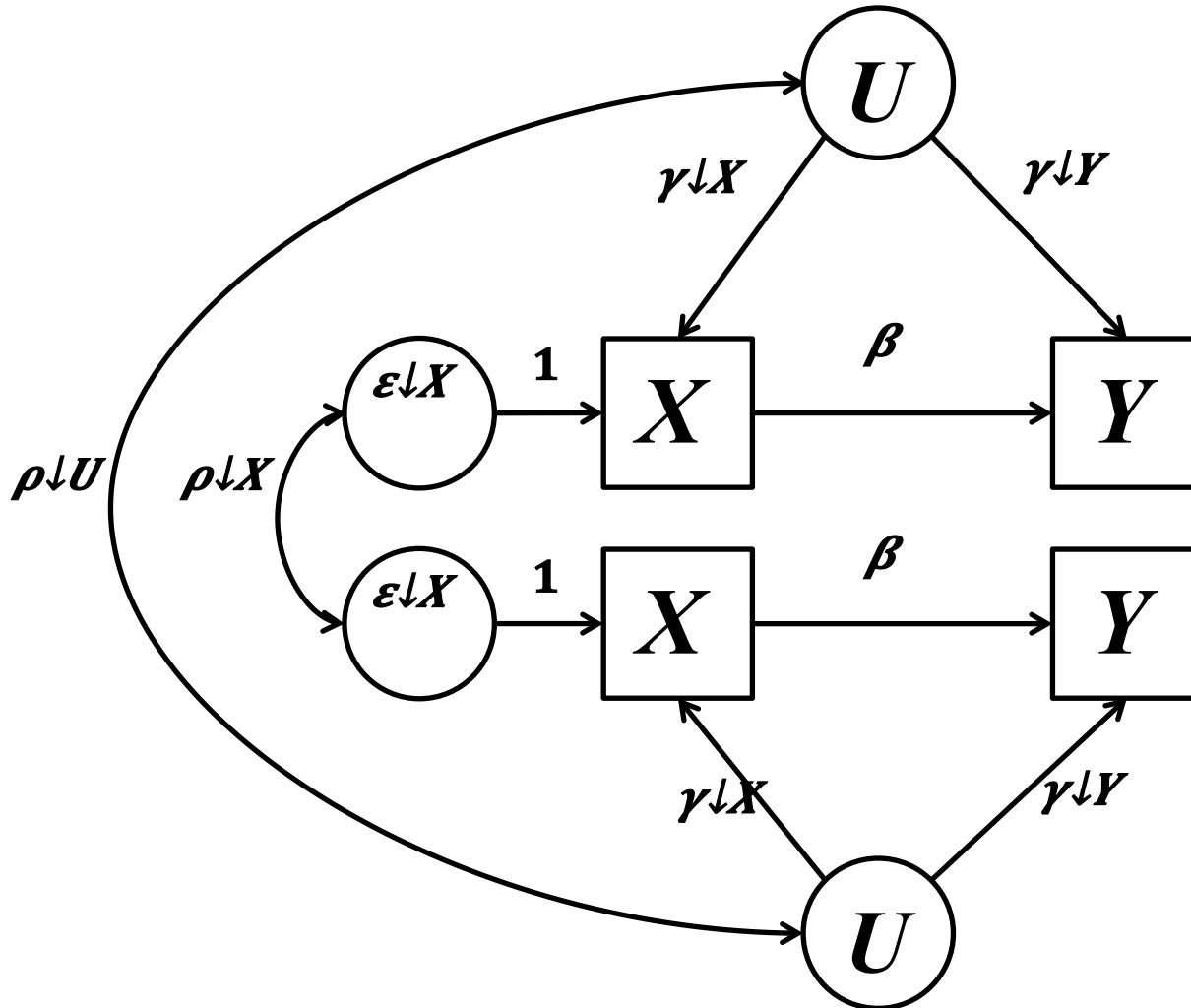


Causality DAG



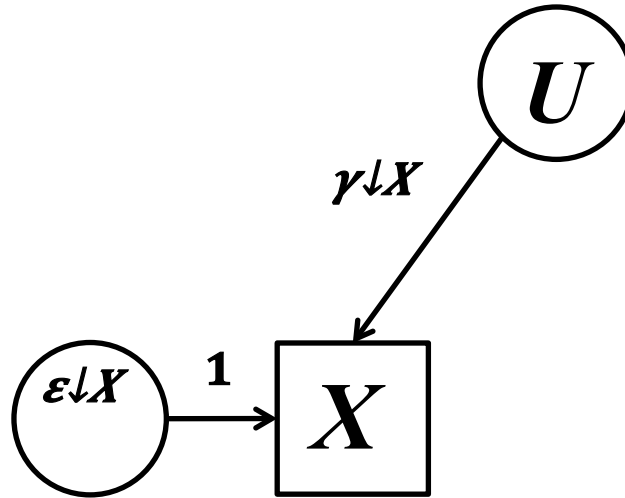
Causality

Path diagram



Causality

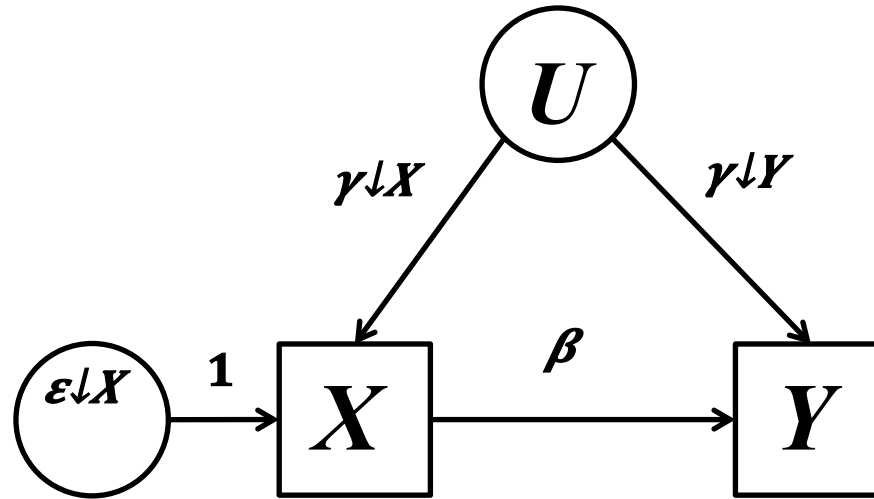
Path diagram



Two processes,
 U and $\epsilon \downarrow X$,
which leads to X

Causality

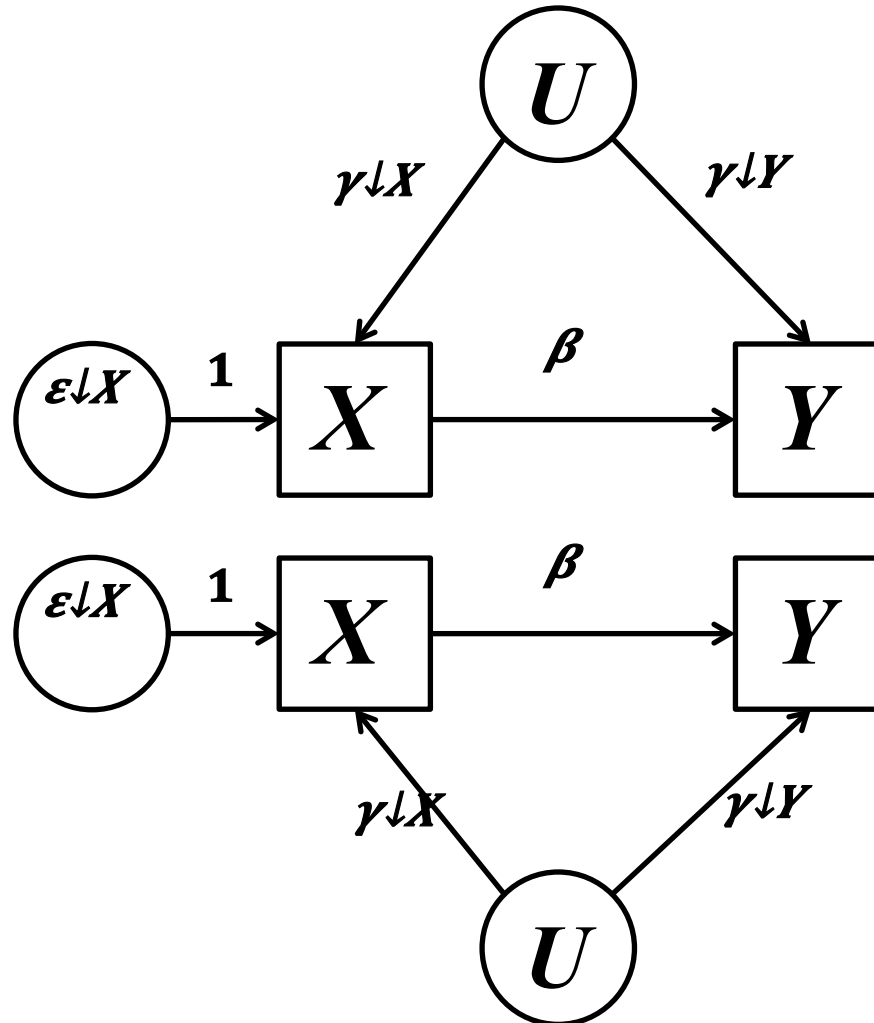
Path diagram



Two processes,
 U and $\varepsilon \downarrow X$,
which leads to X
and Y

Causality

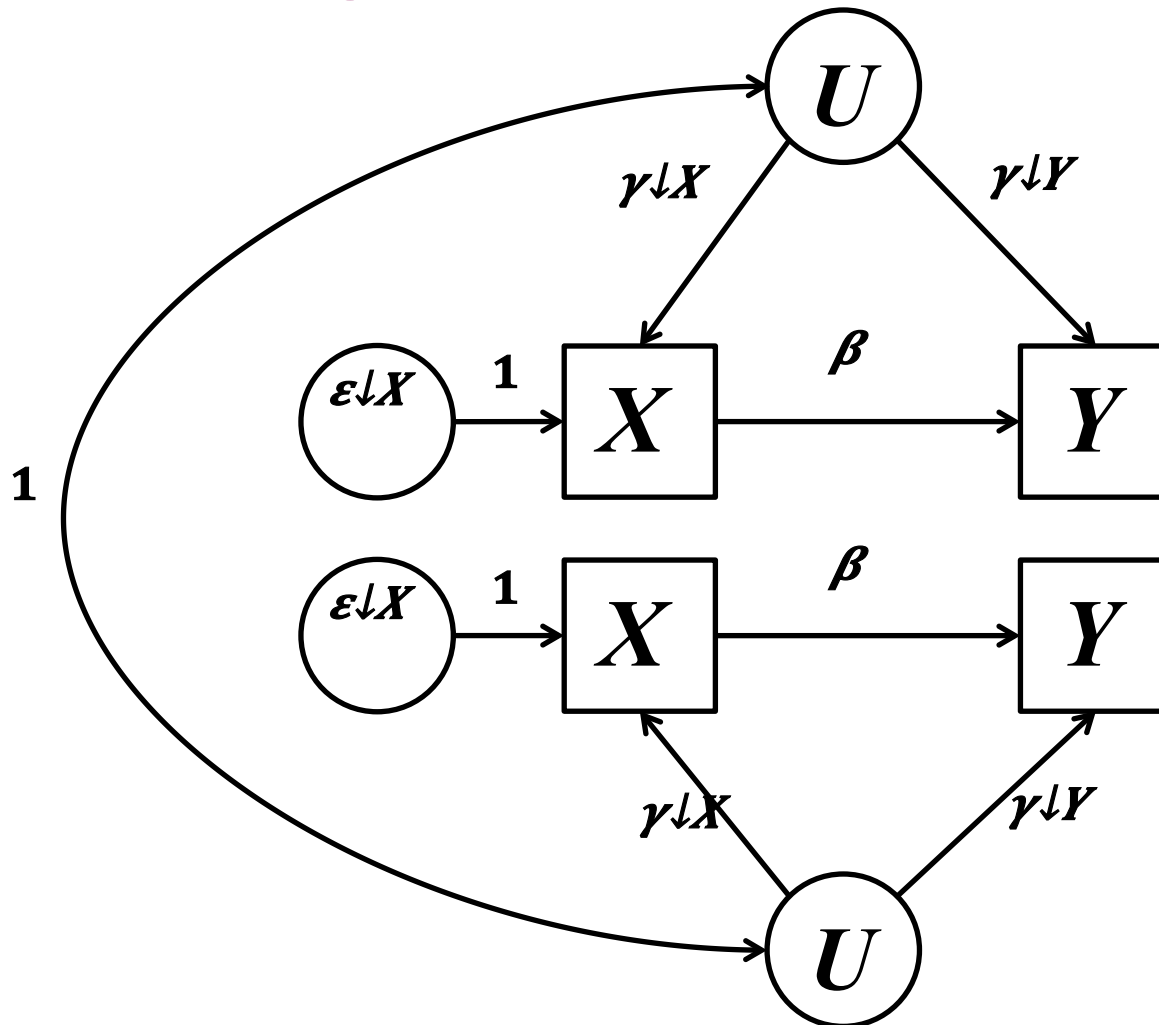
Path diagram



The same process in the other individual in a pair

Causality

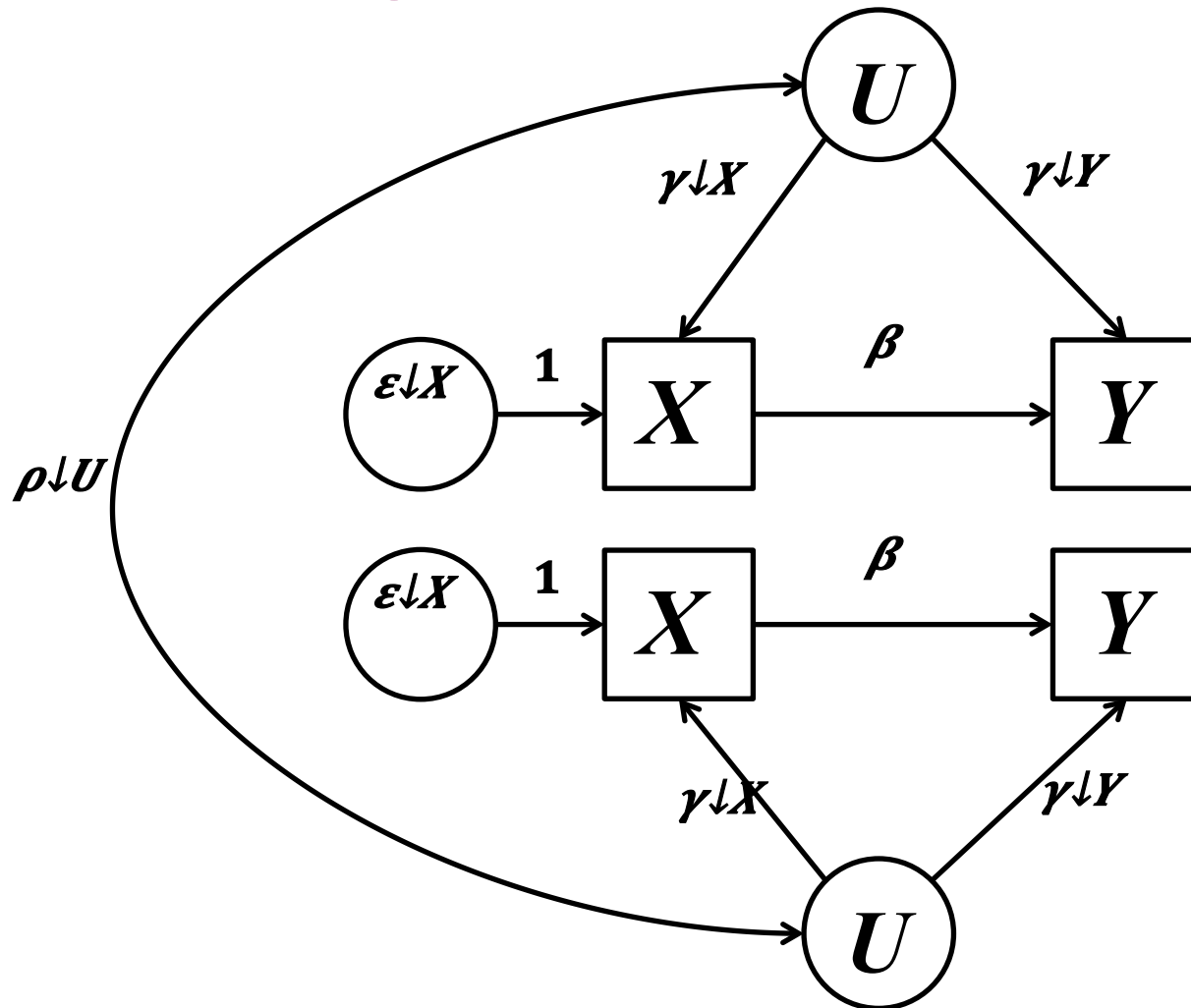
Path diagram



Between-Within
analyses
assumes U to
be perfectly
shared in
pairs...

Causality

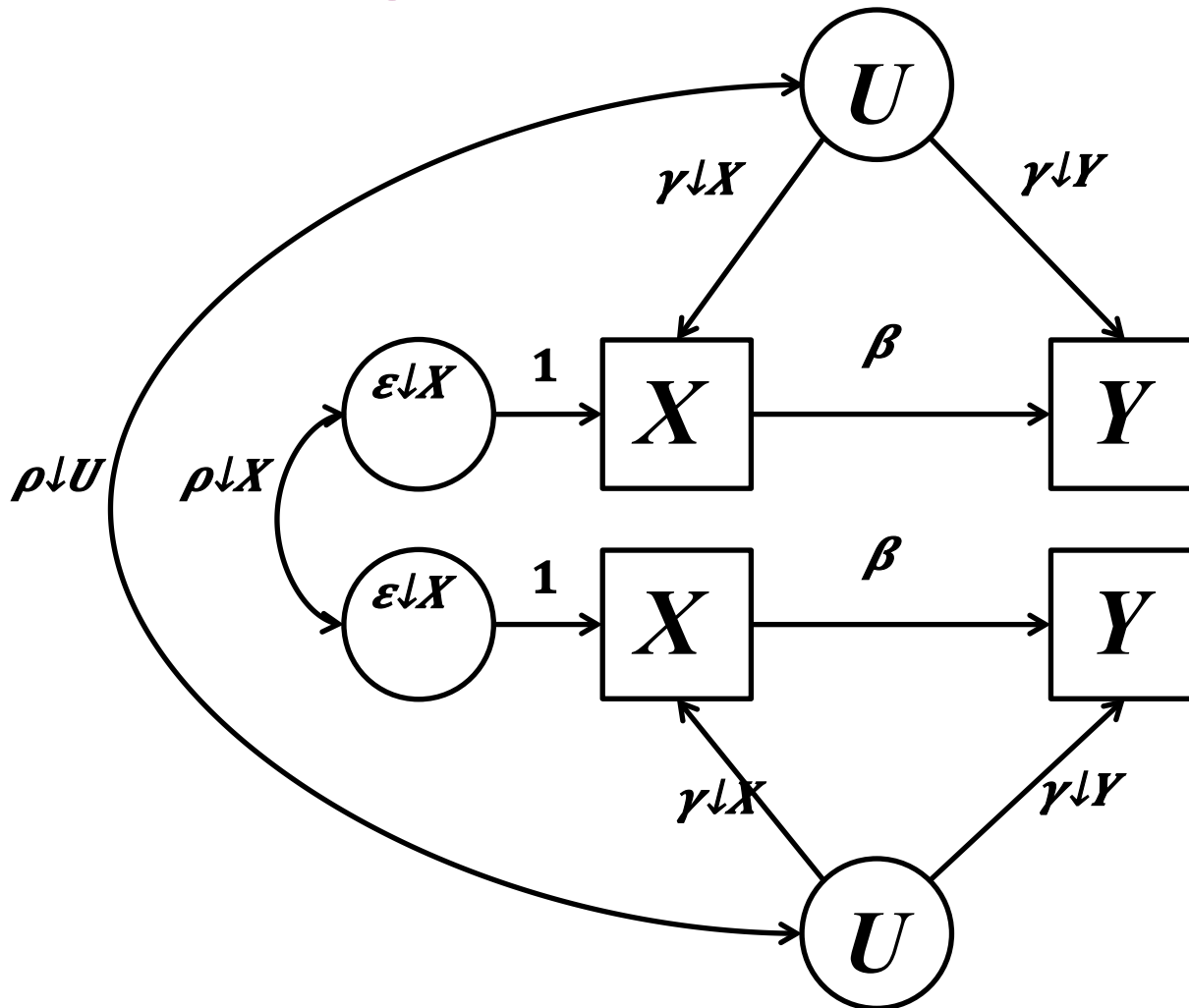
Path diagram



... which
probably isn't
true

Causality

Path diagram



Bias reduction
depends on
 $\rho \downarrow U$ and $\rho \downarrow X$

Between-Within

- Data-generating mechanism

$$X_{ij} = \gamma X_i + U_{ij} + \varepsilon_{Xij}$$

$$g(Y_{ij}) = \beta X_{ij} + \gamma Y_i + U_{ij} + \varepsilon_{Yij},$$

where $Cor(\varepsilon_{Xi1}, \varepsilon_{Xi2}) = \rho_X$ and $Cor(U_{i1}, U_{i2}) = \rho_U$

- Fitted models:

- “Cohort” estimate:

$$g(E(Y_{ij} | X_{ij})) = \mu + \beta_C X_{ij}$$

(1)

- Within estimate:

~~$$g(E(Y_{ij} | X_{ij}, X_{i\cdot})) = \mu + \beta_B X_{i\cdot} + \beta_W$$~~

X_{ij} (2)

Between-Within

- Data-generating mechanism

$$X_{ij} = \gamma_X U_{ij} + \varepsilon_{Xij}$$

$$g(Y_{ij}) = \beta X_{ij} + \gamma_Y U_{ij} + \varepsilon_{Yij},$$

where $Cor(\varepsilon_{Xi1}, \varepsilon_{Xi2}) = \rho_X$ and $Cor(U_{i1}, U_{i2}) = \rho_U$

- Fitted models:

- “Cohort” estimate:

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(1)

- Within estimate:

~~$$g(E(Y_{ij} | X_{ij}, X_{i\cdot})) = \mu + \beta_B X_{i\cdot} + \beta_W$$~~

X_{ij} (2)

Unmeasured confounders

- It can be shown that (if $g(\cdot)$ is the identity link)

$$\beta_{\downarrow C} = \beta + \gamma_{\downarrow Y} \gamma_{\downarrow X} \sigma_{\downarrow U}^2 / \gamma_{\downarrow X}^2 \sigma_{\downarrow U}^2 + \sigma_{\downarrow \varepsilon} \downarrow X \uparrow^2$$

and

$$\beta_{\downarrow W} = \beta + \gamma_{\downarrow Y} \gamma_{\downarrow X} \sigma_{\downarrow U}^2 / \gamma_{\downarrow X}^2 \sigma_{\downarrow U}^2 + \sigma_{\downarrow \varepsilon} \downarrow X \uparrow^2 \frac{1 - \rho_{\downarrow X}}{1 - \rho_{\downarrow U}}$$

Unmeasured confounders

- It can be shown that (if $g(\cdot)$ is the identity link)

$$\beta_C = \beta + \gamma \frac{\sigma_U^2}{\gamma^2 \sigma_X^2 + \sigma_U^2} + \sigma_\epsilon \frac{\sigma_U^2}{\gamma \sigma_X^2 + \sigma_U^2}$$

and

$$\beta_W = \beta + \gamma \frac{\sigma_U^2}{\gamma^2 \sigma_X^2 + \sigma_U^2} + \sigma_\epsilon \frac{\sigma_U^2}{\gamma \sigma_X^2 + \sigma_U^2} \frac{1 - \rho_{XU}}{1 - \rho_{XU}}$$

Bias terms

Unmeasured confounders

- It can be shown that (if $g(\cdot)$ is the identity link)

$$\beta_C = \beta + \gamma \frac{\sigma_U^2}{\sigma_U^2 + \sigma_\varepsilon^2} \frac{\sigma_X^2}{1 - \rho^2} + \sigma_\varepsilon^2 \frac{\rho}{1 - \rho^2}$$

and

$$\beta_W = \beta + \gamma \frac{\sigma_U^2}{\sigma_U^2 + \sigma_\varepsilon^2} \frac{\sigma_X^2}{1 - \rho^2} + \sigma_\varepsilon^2 \frac{\rho}{1 - \rho^2}$$

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and

$$\beta_W = \beta + \gamma \frac{\sigma_U^2}{\sigma_X^2 + \sigma_U^2} + \sigma_{\epsilon}^2 \frac{\sigma_U^2}{\sigma_X^2 + \sigma_U^2} \frac{1 - \rho_{XU}}{1 - \rho_U^2}$$

Bias terms

If $\frac{1 - \rho_{XU}}{1 - \rho_U^2} \rightarrow 1$ then $\beta_W \rightarrow \beta_C$

i.e. when $\rho_{XU} = \rho_U^2$

Unmeasured confounders

- It can be shown that (if $g(\cdot)$ is the identity link)

$$\beta_{\downarrow C} = \beta + \gamma_{\downarrow Y} \gamma_{\downarrow X} \sigma_{\downarrow U}^2 / \gamma_{\downarrow X}^2 \sigma_{\downarrow U}^2 + \sigma_{\downarrow \varepsilon}^2 / \gamma_{\downarrow X}^2$$

and

$$\beta_{\downarrow W} = \beta + \gamma_{\downarrow Y} \gamma_{\downarrow X} \sigma_{\downarrow U}^2 / \gamma_{\downarrow X}^2 \sigma_{\downarrow U}^2 + \sigma_{\downarrow \varepsilon}^2 / \gamma_{\downarrow X}^2 \frac{1 - \rho_{\downarrow X}}{1 - \rho_{\downarrow U}}$$

Bias terms

If $\frac{1 - \rho_{\downarrow X}}{1 - \rho_{\downarrow U}} \rightarrow \infty$ then $\beta_{\downarrow W} \rightarrow \beta$

i.e. when $\rho_{\downarrow U} \rightarrow 1$

Unmeasured confounders

- It can be shown that (if $g(\cdot)$ is the identity link)

$$\beta_{IC} = \beta + \gamma \frac{\sigma_U^2}{\sigma_U^2 + \sigma_{\epsilon|X}^2}$$

and

$$\beta_{IW} = \beta + \gamma \frac{\sigma_U^2}{\sigma_U^2 + \sigma_{\epsilon|X}^2} \frac{1 - \rho_{XU}}{1 - \rho_{XU}^2}$$

Bias terms

When $\rho_{XU} > \rho_{XU}^2$

β_{IW} is **more** biased than β_{IC}

Unmeasured confounders

- It can be shown that (if $g(\cdot)$ is the identity link)

$$\beta_{\downarrow C} = \beta + \gamma_{\downarrow Y} \gamma_{\downarrow X} \sigma_{\downarrow U}^2 / \gamma_{\downarrow X}^2 \sigma_{\downarrow U}^2 + \sigma_{\downarrow \varepsilon} \gamma_{\downarrow X}^2$$

and

$$\beta_{\downarrow W} = \beta + \gamma_{\downarrow Y} \gamma_{\downarrow X} \sigma_{\downarrow U}^2 / \gamma_{\downarrow X}^2 \sigma_{\downarrow U}^2 + \sigma_{\downarrow \varepsilon} \gamma_{\downarrow X}^2 \frac{1 - \rho_{\downarrow X}}{1 - \rho_{\downarrow U}}$$

Bias terms

When $\rho_{\downarrow X} < \rho_{\downarrow U}$

$\beta_{\downarrow W}$ is **less** biased than $\beta_{\downarrow C}$

Unmeasured confounders

When $\rho_{\downarrow X} > \rho_{\downarrow U}$

$\beta_{\downarrow W}$ is **more** biased than $\beta_{\downarrow C}$

When $\rho_{\downarrow X} < \rho_{\downarrow U}$

$\beta_{\downarrow W}$ is **less** biased than $\beta_{\downarrow C}$

Measurement error

- Suppose we imperfectly measure X as
 $X_{ij}^* = X_{ij} + \varepsilon_{ij} M_{ij}$

Measurement error

- Suppose we imperfectly measure X as

$$X_{ij}^* = X_{ij} + \varepsilon_{ij}$$

- Estimates will be **attenuated**, the cohort estimate:

$$\beta_{C^*} = \lambda \beta_C$$

where $\lambda = \text{Var}(X) / (\text{Var}(X) + \text{Var}(\varepsilon))$ is the reliability ($0 \leq \lambda \leq 1$)

Measurement error

- The within estimate will be **further attenuated**

$$\beta_{W*} = \frac{\lambda(\beta\gamma\sigma_U^2(1-\rho_U) + \sigma_{\varepsilon X}^2(1-\rho_X))}{\gamma\sigma_U^2(1-\lambda\rho_U) + \sigma_{\varepsilon X}^2(1-\lambda\rho_X) + \gamma\sigma_U^2(1-\rho_U)}$$

Measurement error

- Suppose $\gamma \downarrow X = 0$ then the estimates will be

$$\beta \downarrow C \uparrow^* = \lambda \beta = \beta(1 - (1 - \lambda))$$

$$\beta \downarrow W \uparrow^* = \beta(1 - 1 - \lambda / 1 - \text{Cor}(X \downarrow i1 \uparrow^*, X \downarrow i2 \uparrow^*))$$

- So even with no confounding the within estimate will be further attenuated

Simulations

Dichotomous

- Simulated model, $Y \sim \text{Binary}$, X and $U \sim \text{MVN}$:

$$OR=2$$

$$\rho_{\downarrow U} = 1$$

$$\rho_{\downarrow X} = 0.5$$

$$\gamma_{\downarrow X} = 0.8$$

$$\gamma_{\downarrow Y} = 0.7$$

$$\lambda = 1$$

$$\rho_{\downarrow U} = 1$$

- Result:

- "Cohort": $OR = 2.7$

- BW: $OR_{\downarrow W} = 2.0$

Simulations

Dichotomous

- Simulated model, $Y \sim \text{Binary}$, X and $U \sim \text{MVN}$:

$$OR=2$$

$$\rho_{U \downarrow X} = 0.2$$

$$\rho_{U \downarrow Y} = 0.5$$

$$\gamma_{U \downarrow X} = 0.8$$

$$\gamma_{U \downarrow Y} = 0.7$$

$$\lambda = 1$$

$$\rho_{U \downarrow Y} < \rho_{U \downarrow X}$$

- Result:

- "Cohort": $OR = 2.7$

- BW: $OR_{\downarrow W} = 2.9$

Simulations

Dichotomous

- Simulated model, $Y \sim \text{Binary}$, X and $U \sim \text{MVN}$:

$$OR=2$$

$$\rho_{U \downarrow X} = 1$$

$$\rho_{U \downarrow Y} = 0.5$$

$$\gamma_{U \downarrow X} = 0.8$$

$$\gamma_{U \downarrow Y} = 0.7$$

$$\lambda = 0.7$$

$$\lambda < 1$$

- Result:

- "Cohort": $OR = 1.9$

- BW: $OR \downarrow W = 1.3$

Simulations

Dichotomous

- Simulated model, $Y \sim \text{Binary}$, X and $U \sim \text{MVN}$:

$$OR=1$$

$$\rho_{U \downarrow X} = 0.8$$

$$\rho_{U \downarrow Y} = 0.5$$

$$\gamma_{U \downarrow X} = 0.8$$

$$\gamma_{U \downarrow Y} = 0.7$$

$$\lambda = 0.9$$

$$OR=1$$

- Result:

- "Cohort": $OR = 1.4$
- BW: $OR_{\downarrow W} = 1.1$

Simulations

- Simulations in software package R
 - eAppendix to article "Sibling Comparison Designs: *Bias From Non-Shared Confounders and Measurement Error*" (Frisell *et al* 2012) e-published ahead of print in *Epidemiology*
 - ... or ask me and I'll send some code to you
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Example:

Cognitive ability and violent offending

Frisell et al 2012

- Criminological research has found lower cognitive ability to increase risk for violent behavior
- ~700,000 Swedish males tested for cognitive ability during conscription were assessed for violent criminal convictions
- ~240,000 full brothers, ~18,000 half brothers reared together and ~25,000 half brothers reared apart

Example:

Cognitive ability and violent offending

RESULTS	Cohort	Within
Full	-0.19	-0.10
Half together	-0.18	-0.13
Half apart	-0.18	-0.16

- The authors investigate different plausible explanations:
 - Increased bias? No, based on previous literature
 - Measurement error?

Reliability (λ)	1.0	0.9	0.8	0.7	0.6	0.5
Full	-0.19	-0.17	-0.14	-0.11	-0.07	-0.01
Half together	-0.18	-0.17	-0.16	-0.14	-0.12	-0.09
Half apart	-0.18	-0.17	-0.16	-0.15	-0.14	-0.12

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Reliability (λ)	1.0	0.9	0.8	0.7	0.6	0.5
Full	-0.19	-0.17	-0.14	-0.11	-0.07	-0.01
Half together	-0.18	-0.17	-0.16	-0.14	-0.12	-0.09
Half apart	-0.18	-0.17	-0.16	-0.15	-0.14	-0.12

Example:

Cognitive ability and violent offending

- Test-retest $\rightarrow \lambda=0.9-0.8$
- The authors conclude that there might be a causal effect
- However the association is partly explained by factors shared increasingly by half brothers apart, half brothers together and full brothers

RESULTS	Within
Full	-0.10
Half together	-0.13
Half apart	-0.16

Reliability (λ)	0.9	0.8
Full	-0.17	-0.14
Half together	-0.17	-0.16
Half apart	-0.17	-0.16

Summary

- The bias can both increase and decrease from within sibling comparisons
- The bias reduction (increase) depends on correlation in confounders vs. exposure
- If there is measurement error the within estimate will be more attenuated than the cohort estimate

Recommendations

- Use sibling comparisons!

... but ...

- Reason about unmeasured confounders and their stability across siblings ($\rho \downarrow U$ vs. $\rho \downarrow X$)
- Is there any information on reasonable values for reliability? In that case use them to see whether that is an alternative explanation

→ Subject matter knowledge

- Frisell Thomas, Öberg Sara, Kuja-Halkola Ralf, Sjölander Arvid. **Sibling Comparison Designs: Bias From Non-Shared Confounders and Measurement Error.** *Epub ahead of print Epidemiology 2012*

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References

- Frisell Thomas, Öberg Sara, Kuja-Halkola Ralf, Sjölander Arvid. Sibling Comparison Designs: Bias From Non-Shared Confounders and Measurement Error. *Epub ahead of print Epidemiology* 2012.
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- Sjölander Arvid, Johansson Anna L.V. , Lundholm Cecilia , Altman Daniel, Almqvist Catarina, Pawitan Yudi. Analysis of 1:1 matched cohort studies and twin studies, with binary exposures and binary outcomes. *Accepted for publication in Statistical Science* 2012.
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